# CRITIQUE OF HIRSCH'S CITATION INDEX: A COMBINATORIAL FERMI PROBLEM 

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## 1. Introduction

1.1. Overview. In 2005, physicist J. E. Hirsch Hi05] proposed the $h$-index to measure the quality of a researcher's output. This metric is the largest integer $n$ such that the person has $n$ papers with at least $n$ citations each, and all other papers have weakly less than $n$ citations. Although the original focus of loc. cit. was on physicists, the $h$-index is now widely popular. For example, Google Scholar and the Web of Science highlight the $h$-index, among other metrics such as total citation count, in their profile summaries.

An enticing point made in loc. cit. is that the $h$-index is an easy and useful supplement to a citation count ( $N_{\text {citations }}$ ), since the latter metric may be skewed by a small number of highly cited papers or textbooks. In Hirsch's words:
"I argue that two individuals with similar $h$ s are comparable in terms of their overall scientific impact, even if their total number of papers or their total number of citations is very different. Conversely, comparing two individuals (of the same scientific age) with a similar number of total papers or of total citation count and very different $h$ values, the one with the higher $h$ is likely to be the more accomplished scientist."
It seems to us that users might tend to eyeball differences of $h \mathrm{~s}$ and citation counts among individuals during their assessments. Instead, one desires a quantitative baseline for what "comparable", "very different" and "similar" actually mean. Now, while this would appear to be a matter for statisticians, we show how textbook combinatorics sheds some light on the relationship between the $h$-index and $N_{\text {citations }}$. We present a simple model that raises specific concerns about potential misuses of the $h$-index.

To begin, think of the list of a researcher's citations per paper in decreasing order $\lambda=$ ( $\lambda_{1} \geq \lambda_{2} \geq \cdots \geq \lambda_{N_{\text {papers }}}$ ) as a partition of size $N_{\text {citations }}$. Graphically, $\lambda$ is identified with its Young diagram. For example, $\lambda=(5,3,1,0) \leftrightarrow \underset{\bullet \bullet}{\bullet} \bullet \square$.

A combinatorialist will recognize that the $h$-index of $\lambda$ is the side-length of the Durfee square (marked using $\bullet$ 's above): this is the largest $h \times h$ square that fits in $\lambda$. This simple observation is nothing new, and appears in both the bibliometric and combinatorial literature, see, e.g., [AnHaKi09, FlSe09]. In particular, since the Young diagram of size $N_{\text {citations }}$ with maximum $h$-index is roughly a square, we see graphically that $0 \leq h \leq\left\lfloor\sqrt{N_{\text {citations }}}\right\rfloor$.

Next, consider the following question:
Given $N_{\text {citations }}$, what is the estimated range of $h$ ?
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Taking only $N_{\text {citations }}$ as input hardly seems like sufficient information to obtain a meaningful answer. It is exactly for this reason that we call the question a combinatorial Fermi problem, by analogy with usual Fermi problems; see Section 2.

Since we assume no prior knowledge, consider each citation profile in an unbiased manner. That is, each partition of $N_{\text {citations }}$ is chosen with equal probability. In fact, there is a beautiful theory concerning the asymptotics of these uniform random partitions. This was largely developed by A. Vershik and his collaborators; see, e.g., the survey [Su10].

Actually, we are interested in "low" (practical) values of $N_{\text {citations }}$ where not all asymptotic results are exactly relevant. Instead, we use generating series and modern desktop computation to calculate the probability that a random $\lambda$ has Durfee square size $h$. More specifically, we obtain Table 1 below using the Euler-Gauss identity for partitions:

$$
\begin{equation*}
\prod_{i=1}^{\infty} \frac{1}{1-x^{i}}=1+\sum_{k \geq 1} \frac{x^{k^{2}}}{\prod_{j=1}^{k}\left(1-x^{j}\right)^{2}} \tag{1}
\end{equation*}
$$

The proof of (1) via Durfee squares is regularly taught to undergraduate combinatorics students; it is recapitulated in Section 3. The pedagogical aims of this note are elaborated upon in both Sections 2 and 3 .

| $N_{\text {citations }}$ | 50 | 100 | 200 | 300 | 400 | 500 | 750 | 1000 | 1250 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Interval for $h$ | $[2,5]$ | $[3,7]$ | $[5,9]$ | $[7,11]$ | $[8,13]$ | $[9,14]$ | $[11,17]$ | $[13,20]$ | $[15,22]$ |


| 1500 | 1750 | 2000 | 2500 | 3000 | 3500 | 4000 | 4500 | 5000 | 5500 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $[17,24]$ | $[18,26]$ | $[20,28]$ | $[22,31]$ | $[25,34]$ | $[27,36]$ | $[29,39]$ | $[31,41]$ | $[34,43]$ | $[35,45]$ |


| 6000 | 6500 | 7000 | 7500 | 8000 | 9000 | 10000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $[36,47]$ | $[37,49]$ | $[39,51]$ | $[40,52]$ | $[42,54]$ | $[44,57]$ | $[47,60]$ |

TABLE 1. Confidence intervals for $h$-index

The asymptotic result we use, due to E. R. Canfield-S. Corteel-C. D Savage [CaCoSa98], gives the mode size of the Durfee square when $N_{\text {citations }} \rightarrow \infty$. Since their formula is in line with our computations, even for low $N_{\text {citations, }}$, we reinterpret their work as the

$$
\text { rule of thumb for } h \text {-index: } \quad h=\frac{\sqrt{6} \log 2}{\pi} \sqrt{N_{\text {citations }}} \approx 0.54 \sqrt{N_{\text {citations }}} .
$$

The focus of this paper is on mathematicians. For the vast majority of those tested, the actual $h$-index computed using Mathscinet or Google scholar falls into the confidence intervals. Moreover, we found that the rule of thumb is fairly accurate for pure mathematicians. For example, Table 2 shows this for post-1998 Fields medalists..$^{\top}$

In [Hi05] it was indicated that the $h$-index has predictive value for winning the Nobel prize. However, the relation of $h$ index to the Fields medal is, in our opinion, unclear. A number of the medalists' $h$ values below are shared (or exceeded) by non-contenders of

[^0]similar academic age, or with those who have the similar citation counts. Perhaps, this is reflects a cultural difference between the mathematics and the scientific communities.

In Section 4, we analyze mathematicians in the National Academy of Sciences, where we show the correlation between the rule of thumb and actual $h$-indices is $R=0.94$. After removing book citations, $R=0.95$. We also discuss Abel prize winners and associate professors at three research universities.

Ultimately, the reader is encouraged to do checks of the estimates themselves.

| Medalist | Award year | $N_{\text {citations }}$ | $h$ | Rule of thumb est. | Confidence interval |
| :---: | :---: | :---: | :---: | :---: | :---: |
| T. Gowers | 1998 | 1012 | 15 | 17.2 | $[13,20]$ |
| R. Borcherds | 1998 | 1062 | 14 | 17.6 | $[14,21]$ |
| C. McMullen | 1998 | 1738 | 25 | 22.5 | $[18,26]$ |
| M. Kontsevich | 1998 | 2609 | 23 | 27.6 | $[22,32]$ |
| L. Lafforgue | 2002 | 133 | 5 | 6.2 | $[4,8]$ |
| V. Voevodsky | 2002 | 1382 | 20 | 20.0 | $[16,23]$ |
| G. Perelman | 2006 | 362 | 8 | 10.0 | $[7,12]$ |
| W. Werner | 2006 | 1130 | 19 | 18.2 | $[14,21]$ |
| A. Okounkov | 2006 | 1677 | 24 | 22.1 | $[18,25]$ |
| T. Tao | 2006 | 6730 | 40 | 44.3 | $[38,51]$ |
| C. Ngô | 2010 | 228 | 9 | 8.2 | $[9,14]$ |
| E. Lindenstrauss | 2010 | 490 | 12 | 12.0 | $[9,15]$ |
| S. Smirnov | 2010 | 521 | 12 | 12.3 | $[24,33]$ |
| C. Villani | 2010 | 2931 | 25 | 29.2 |  |

TABLE 2. Fields medalists $1998-2010$

We discuss three implications/possible applications of our analysis.
1.2. Comparing $h$ 's when $N_{\text {citations }}$ 's are very different. It is understood that $h$-index usually grows with $N_{\text {citations. }}$. However, when are citation counts so different that comparing $h^{\prime}$ s is uninformative? For example, $h_{\text {Tao }}=40$ ( 6,730 citations) while $h_{\text {Okounkov }}=24$ ( 1,677 citations). The model asserts the probability of $h_{\text {Okounkov }} \geq 32$ is less than 1 in 10 million. Note the Math genealogy project has fewer than 200, 000 mathematicians.

These orders of magnitude predict that no mathematician with 1,677 citations has an $h$-index of 32, even though technically it can be as high as 40 . Similarly, one predicts the rarity of pure mathematicians with these citations having "similar" $h$-indices (in the pedestrian sense). This is relevant when comparing (sub)disciplines with vastly different typical citation counts. We have a theoretical caution about "eyeballing".
1.3. The rule of thumb and the highly cited. The model suggests the theoretical behavior of the $h$-index for highly-cited scholars. The extent to which these predictions hold true is informative. This is true not only for individuals, but for entire fields as well.

Actually, Hirsch defined a proportionality constant $a$ by $N_{\text {citations }}=a h^{2}$ and remarked, "I find empirically that $a$ ranges between 3 and 5 ." This asserts $h$ is between $\sqrt{1 / 5} \approx 0.45$ and $\sqrt{1 / 3} \approx 0.58$ times $\sqrt{N_{\text {citations }}}$.

One can begin to try to understand the similarity between Hirsch's empirical upper bound and the rule of thumb. A conjecture of E. R. Canfield (private communication, see

Section 3) asserts concentration around the mode Durfee square. Thus, theoretically, one expects the rule of thumb to be nearly correct for $N_{\text {citations }}$ large.

Alas, this is empirically not true, even for pure mathematicians. However, we observe something related: $0.54 \sqrt{N_{\text {citations }}}$ is higher than the actual $h$ for almost every very highly cited ( $N_{\text {citations }}>10,000$ ) scholar in mathematics, physics, computer science and statistics (among others) we considered. On the rare occasion this fails, the estimate is only beat by a small percentage ( $<5 \%$ ). The drift in the other direction is often quite large ( $50 \%$ or more is not unusual in certain areas of engineering or biology).

Near equality occurs among Abel prize winners. We also considered all prominent physicists highlighted in [Hi05] (except Cohen and Anderson, due to name conflation in Web of Science). The guess is always an upper bound (on average $14-20 \%$ too high). Near equality is met by D. J. Scalapino (25, 881 citations; 1.00), C. Vafa (22, 902 citations; $0.99)$, J. N. Bahcall (27, 635 citations; 0.98 ); we have given the ratio $\frac{\text { true } h}{\text { estimated } h}$.

One reason for highly cited people to have lower than expected $h$-index is that they tend to have highly cited textbooks. Also, famous academics often run into the "Matthew effect" (e.g., gratuitous citations of their most well-known articles or books).
1.4. Anomalous $h$-indices. More generally, our estimates give a way to flag anomalous $h$-indices of active researchers, i.e., those that are far outside the confidence interval, or, e.g., those for which the rule of thumb is especially inaccurate.

To see what effect book citations has on our estimates, consider the combinatorialist R. P. Stanley. Since Stanley has 6,510 citations, we estimate his index as 43.6. However, $h_{\text {Stanley }}=35$, a $20 \%$ error. Now, 3,237 of his citations come from textbooks. Subtracting these, one estimates his $h$-index as 30.9 while his revised actual $h$-index is 32 , only a $4 \%$ error. This kind of phenomenon was not uncommon; see the appendix.

For another example, consider T. Tao's Google scholar profile. Since he has 30, 053 citations, the rule of thumb predicts his $h$-index is 93.6. This is far from his actual $h$-index of 65. Now, his top five citations (joint with E. Candes on compressed sensing) are applied. Removing the papers on this topic leaves 13,942 citations. His new estimate is therefore 63.7 and his revised $h$-index is 61 .

In many cases we have looked at, once the "skewing" feature of the scholar's profile is removed, the remainder of their profile agrees with the rule of thumb.
1.5. Conclusions and summary. Whether it be Fields medalists, Abel prize laureates, job, promotion or grant candidates, clearly, the quality of a researcher cannot be fully measured by numerics. However, in reality, the $h$-index is used, formally or informally, for comparisons. This paper attempts to provide a theoretical and testable framework to quantitatively understand the limits of such evaluations. For mathematicians, the accuracy of the rule of thumb suggests that the differences of $h$ index between two mathematicians is strongly influenced by their respective citation counts.

While discussion of celebrated mathematicians and their statistics makes for fun coffee shop chatter, a serious way that $h$-index comes up in faculty meetings concerns relatively junior mathematicians. Consider a scholar $A$ with 100 citations and $h$ index of 6 and a scholar $B$ who has 50 citations and an $h$-index of 4 . Such numbers are not atypical of math assistant professors going up for tenure. Our model predicts $h_{A}$ to be a little bigger than $h_{B}$. Can one really discern what portion of $h_{A}-h_{B}$ is a signal of quality?

The problem becomes larger when $A$ and $B$ are in different subject areas. Citations for major works in applied areas tend to have many more citations than in mathematics. In experimental fields, papers may have many coauthors. Since $h$-index does not account for authorship order, this tends to affect our estimates for such subjects.

Pure mathematicians have comparatively fewer coauthors, papers and citations. It is not uncommon for, e.g., solutions to longstanding open problems, to have relatively few citations. Thus an explanatory model for pure mathematicians has basic reasons for being divergent for some other fields. Yet, if this is the case, can the $h$-index really be used universally? This gives us a theoretical reason to question whether one can make simple comparisons across fields, even after a rescaling, as has been suggested in [IgPe07].

## 2. Combinatorial Fermi Problems

2.1. Usual Fermi problems. Fermi problems are so-named after E. Fermi, whose ability to obtain good approximate quantitative answers with little data available is legendary. As an illustration, we use the following example [Co]:

## How many McDonalds operate in the United States?

There are 10 McDonalds in Champaign county, which has a population of about 200, 000 . Assume the number of McDonalds scales with population. Since the population of the United States is 300 million, a "back-of-the-envelope" calculation estimates the number of McDonalds at 15,000 . The actual answer, as of 2012, is 14,157 .

Using a simplified assumption like the italicized one above is a feature of a Fermi problem. Clearly, the uniform assumption made is not really correct. However, the focus is on good, fast approximations when more careful answers are either too time consuming to determine, or maybe even impossible to carry out. The approximation can then be used to guide further work to determine more accurate/better justified answers.

Now, although the estimate is rather close to the actual number, when the estimate is not good, the result is even more interesting, as it helps identify a truly faulty assumption. For instance, analogous analysis predicts that the number of Whole Foods in the United States is 0 . Apparently, the presence of that company does not scale by population.

Fermi problems/back-of-the-envelope calculations are a standard part of a physics or engineering education. They are of theoretical value in the construction mathematical models, and of "real world" value in professions such as management consulting. However, perhaps because the concept is intrinsically non-rigorous, it is not typically part of a (pure) mathematics curriculum. Specifically, this is true for enumerative combinatorics, even though the subject's purpose is to count the number of certain objects - which in the author's experience, many students hope has non-theoretical applicability.
2.2. A combinatorial analogue. By analogy we define a combinatorial Fermi problem:

Fix $\epsilon>0$. Let $S$ be a finite set of combinatorial objects and $\omega: S \rightarrow \mathbb{Z}_{\geq 0}$ be a statistic on $S$. Then we estimate the value of $\omega$ on any element to be the confidence interval $[a, b]$ where the uniform probability of picking an element of $S$ outside of this range has probability $<\epsilon$.
By definition, the (ordinary) generating series for the combinatorial problem $(S, \omega)$ is defined by $G_{(S, \omega)}(x)=\sum_{s \in S} x^{\omega(s)}$. For any $k, \#\{s \in S: \omega(s)=k\}=\left[x^{k}\right] G_{(S, \omega)}(x)$, i.e., the
coefficient of $x^{k}$ in $G_{(S, \omega)}(x)$. Usually, textbook work involves extracting the coefficient using formulae valid in the ring of formal power series. However, what is often not emphasized in class is that this coefficient, and $\# S$ itself can be rapidly extracted using a computer algebra system, allowing for a quick determination of the range $[a, b]$. Since the computer does the work, this is our analogue of a "back-of-the envelope" calculation.

For "reasonable" values of $\epsilon$ (such as $\epsilon=2 \%$ ), often the range $[a, b]$ is rather tight. In those cases, there may be a theorem of asymptotic concentration near a "typical" object. However, even if such theorems are known, this does not solve the finite problem.

The use of the uniform distribution is a quick way to exactly obtain estimates that can be compared with empirical data. Ultimately, it invites the user to consider other probability distributions and more sophisticated statistical analysis (just as one should with the McDonald's example), using e.g., Markov Chain Monte Carlo techniques.

We mention another combinatorial Fermi problem we have considered elsewhere: the count of the number of indigeneous language families in the Americas [Yo13]. That is a situation where essentially there is no way to know with great certainty the true answer.

## 3. The Euler-Gauss identity and its application to the h-Indices

We apply the perspective of Section 2 to the $h$-index question, where $S=\operatorname{Par}(n)$ and $\omega: S \rightarrow \mathbb{Z}_{\geq 0}$ is the size of a partition's Durfee square. If Par is the set of all partitions and $\sigma: \operatorname{Par} \rightarrow \mathbb{Z}_{\geq 0}$ returns the size of a partition, then the generating series for ( $\mathrm{Par}, \sigma$ ) is $P(x)=\prod_{i=1}^{\infty} \frac{1}{1-x^{i}}$. That is, $\# \operatorname{Par}(n)=\left[x^{N}\right] P(x)$. A sample textbook reference is [Br10].

Recall the Euler-Gauss identity (1) from the introduction. The well-known combinatorial proof is that every Young diagram $\lambda$ bijectively decomposes into a triple ( $D_{\lambda}, R_{\lambda}, B_{\lambda}$ ) where $D_{\lambda}$ is a $k \times k$ square, $R_{\lambda}$ is a Young diagram with at most $k$ rows and $B_{\lambda}$ is a partition with at most $k$ columns. That is, $D_{\lambda}$ is the Durfee square, $R_{\lambda}$ be the shape to the right of the square and $B_{\lambda}$ to be the shape below it. For example:

$$
\lambda=\stackrel{\bullet \bullet}{\bullet \bullet} \square \mapsto(\square, \square \square \square, \square)=\left(D_{\lambda}, R_{\lambda}, B_{\lambda}\right) .
$$

The generating series for partitions with at most $k$ columns is directly $\prod_{j=1}^{k} \frac{1}{1-x^{j}}$. Since conjugation (the "transpose") of shape with at most $k$ rows returns a shape with at most $k$ columns, it follows that the generating series for shapes of the first kind is also $\prod_{j=1}^{k} \frac{1}{1-x^{j}}$.

From this argument, we see that the generating series for Young diagrams with Durfee square of size $k$ is $x^{k^{2}} \prod_{j=1}^{k}\left(1-x^{j}\right)^{-2}$. We compute for fixed $h, N_{\text {citations }}$ :

$$
\operatorname{Prob}\left(\lambda:|\lambda|=N_{\text {citations }}, \text { Durfee square of size } k\right)=\frac{\left[x^{\left.N_{\text {citations }}\right]} x^{k^{2}} \prod_{j=1}^{k}\left(1-x^{j}\right)^{-2}\right.}{\operatorname{Par}\left(N_{\text {citations }}\right)} .
$$

Often textbook analysis ends at the derivation of (1). In a classroom, using a computer to Taylor expand $\sum_{k=a}^{b} x^{k^{2}} \prod_{j=1}^{k}\left(1-x^{j}\right)^{-2}$, and comparing the coefficients with the known partition numbers allows the instructor to "physically" demonstrate the identity to the student. Varying $a$ and $b$ shows what range of Durfee square sizes are, e.g., $98 \%$ likely to
occur for partitions of that size. Interpreted in terms of our $h$-index problem, these same computations are what gives us Table 1. ${ }^{2}$

As we state in Section 1, the work of [CaCoSa98] shows the mode Durfree square size is $\approx 0.54 \sqrt{N_{\text {citations }}}$. E. R. Canfield's concentration conjecture states that for each $\epsilon>0$,

$$
\begin{equation*}
\lim _{N_{\text {citations }} \rightarrow \infty} \frac{\text { \# partitions with }(1-\epsilon) \mu<h<(1+\epsilon) \mu}{\# \operatorname{Par}\left(N_{\text {citations }}\right)} \rightarrow 1, \tag{2}
\end{equation*}
$$

where $\mu=\frac{\sqrt{6} \log 2}{\pi} \sqrt{N_{\text {citations }}}$. This is consistent with Table 1. Further discussion may appear elsewhere. Also, one would like to examine other distributions on Young diagrams, such as the Plancherel measure, which assigns the shape $\lambda$ the probability $\left(f^{\lambda}\right)^{2} /|\lambda|$ ! where $f^{\lambda}$ is the number of standard Young tableaux of shape $\lambda$.

## 4. FURTHER COMPARISONS WITH EMPIRICAL DATA

4.1. The National Academy of Science. We compared our rule of thumb against all 120 mathematicians of the National Academy of Sciences (see the appendix). The correlation coefficient is $R=0.93$. After removing books (as identified in Mathscinet), $R=0.95$. A serious concern is that many pre-2000 citations are not tabulated in Mathscinet. Nevertheless, in our opinion, the results are still informative. See the comments in Section 4.4.


Figure 1. Rule of thumb ( $x$-axis) versus acutal $h$ 's ( $y$-axis) for Mathematics members of the National Academy of Sciences

[^1]

FIGURE 2. Rule of thumb ( $x$-axis) versus actual $h$ 's ( $y$-axis) for Mathematics members of the National Academy of Sciences (with books removed)
4.2. Abel prize winners. Perhaps a closer analogy to the Nobel prize than the Fields medal is the Abel prize, since the latter does not have an age-limit. The fit with the estimated intervals remains decent; the concern about pre-2000 citations remains.

| Laureate | Award year | $N_{\text {citations }}$ | $h$ | rule of thumb est. | Estimated range |
| :---: | :---: | :---: | :---: | :---: | :---: |
| J. P. Serre | 2003 | 10119 | 53 | 54.3 | $[47,60]$ |
| I. Singer | 2004 | 2982 | 28 | 29.5 | $[24,34]$ |
| M. Atiyah | 2004 | 6564 | 40 | 43.7 | $[37,49]$ |
| P. Lax | 2005 | 4601 | 30 | 36.6 | $[31,42]$ |
| L. Carleson | 2006 | 1980 | 18 | 24.0 | $[19,28]$ |
| S. R. S. Varadhan | 2007 | 2894 | 28 | 29.0 | $[24,33]$ |
| J. Thompson | 2008 | 789 | 14 | 15.2 | $[11,18]$ |
| J. Tits | 2008 | 3463 | 28 | 31.8 | $[27,36]$ |
| M. Gromov | 2009 | 7671 | 41 | 47.3 | $[40,54]$ |
| J. Tate | 2010 | 2979 | 30 | 29.5 | $[24,34]$ |
| J. Milnor | 2011 | 7856 | 48 | 47.9 | $[41,54]$ |
| E. Szemerédi | 2012 | 2536 | 26 | 27.2 | $[22,31]$ |
| P. Deligne | 2013 | 6567 | 36 | 43.8 | $[37,50]$ |

TABLE 3. Abel prize recipients
4.3. Associate Professors. Finally, in Table 4 we considered all mathematics associate professors at three research universities. Of the 32 professors, all but five have their $h$ index in the estimated range, and all are at most one unit outside this range.

|  | $N_{\text {citations }}$ | $h$ | rule of thumb est. | estimated range |
| :---: | :---: | :---: | :---: | :---: |
| Department A |  |  |  |  |
| A1 | 19 | 3 | 2.4 | [1, 3] |
| A2 | 80 | 6 | 4.8 | 3, 6] |
| A3 | 113 | 6 | 5.7 | [4, 7] |
| A4 | 130 | 4 | 6.1 | [4, 8] |
| A5 | 202 | 6 | 7.7 | 5,10] |
| A6 | 511 | 11 | 12.2 | 9,15] |
| Department B B1 | 30 | 3 | 3.0 | [1, 4] |
| B2 | 35 | 4 | 3.2 | [2, 4] |
| B3 | 56 | 4 | 4.0 | [2,5] |
| B4 | 56 | 5 | 4.0 | [2,5] |
| B5 | 63 | 5 | 4.3 | 3, 5] |
| B6 | 63 | 6 | 4.3 | 3, 5] |
| B7 | 78 | 3 | 4.8 | 3,6] |
| B8 | 84 | 5 | 4.9 | [3, 6] |
| B9 | 88 | 7 | 5.1 | [3, 6] |
| B10 | 122 | 8 | 6.0 | 4, 7] |
| B11 | 126 | 7 | 6.1 | [4, 7] |
| B12 | 133 | 6 | 6.2 | 4, 8] |
| B13 | 133 | 7 | 6.2 | 4, 8] |
| B14 | 150 | 8 | 6.6 | [4, 8] |
| B15 | 163 | 7 | 6.9 | [5, 8] |
| B16 | 228 | 10 | 8.1 | 5,10] |
| $\begin{gathered} \text { Department C } \\ \text { C1 } \end{gathered}$ | 10 | 2 | 1.7 | [1, 2] |
| C2 | 11 | 2 | 1.8 | [1, 2] |
| C3 | 25 | 3 | 2.7 | [1,3] |
| C4 | 54 | 4 | 4.0 | [2,5] |
| C5 | 64 | 5 | 4.3 | 3, 5] |
| C6 | 64 | 5 | 4.3 | [3, 5] |
| C7 | 67 | 6 | 4.4 | 3, 5] |
| C8 | 104 | 6 | 5.5 | [4, 7] |
| C9 | 144 | 8 | 6.5 | [4, 8] |
| C10 | 269 | 5 | 8.9 | [6,11] |

TABLE 4. Associate professors at three research universities
4.4. Further study. It seems to us that the simple model presented describes one force governing $h$-index. However, other forces/sources of noise are at play, depending on the field or even the fame of the scholar. Future work seeks to better understand this quantitatively, as one works towards more precise models.

The loss of pre-2000 citations in Mathscinet is significant to how we interpret the results for the National Academy members/Abel prize winners. For example, the rough agreement with the rule of thumb might only reflect an "equilibrium state" that arises years
after major results have been published. This concern is partly are allayed by the similar agreement for recent Fields medalists (Table 2). However, as Mathscinet reaches further back in tabulating citations, one would try to quantify these effects. In the meantime, use of Mathscinet has practical justification since in promotion and grant decision cases, recent productivity is important. So for these purposes, post-2000 data is mostly sufficient.

As a further cross-check, we used the rule of thumb for a broad range of fields using Google Scholar. For scholars with a moderate number of citations, the agreement is often similarly good. Also the rule is an upper bound for the vast majority of highly cited scholars (but as we have said earlier, much less accurate in some fields). However, these checks have an obvious bias as they only consider people who have set up a profile, so we do not formally present these results here.

We propose using the rule of thumb and the confidence intervals as a basis for a systematic study. We suggest that the rule of thumb reflects an "ideal scholar". (This terminology is an allusion to "ideal gas" in statistical mechanics. Indeed, a more conventional use of random partitions concerns the study of Boltzmann statistics on a one-dimensional lattice fermion gas.) Divergence from this ideal is a result of "anomalies". For a choice of field, can one statistically distinguish, on quantifiable grounds, scholars who are close to the rule of thumb (in the sense of confidence intervals) from those who are far from it?

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Table 5. (Appendix) Current National Academy of Sciences Members (Mathematics)

| Member | $N_{\text {citations }}$ | Rule of thumb est. | $h$ | non-books only | revised est. | revised $h$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| G. Andrews | 4866 | 37.7 | 28 | 2579 | 27.4 | 24 |
| M. Artin | 2326 | 26 | 26 | 2097 | 24.7 | 24 |
| M. Aschbacher | 1386 | 20 | 17 | 911 | 16.3 | 13 |
| R. Askey | 2480 | 26.9 | 17 | 1235 | 19.0 | 16 |
| M. Atiyah | 6564 | 43.7 | 40 | 5390 | 39.6 | 38 |
| H. Bass | 2472 | 26.8 | 22 | 1869 | 23.3 | 22 |
| E. Berlekamp | 764 | 14.9 | 12 | 363 | 10.3 | 10 |
| J. Bernstein | 2597 | 27.5 | 22 | 2484 | 26.9 | 21 |
| S. Bloch | 1497 | 20.9 | 20 | 1363 | 19.9 | 18 |
| E. Bombieri | 1746 | 22.6 | 23 | 1608 | 21.7 | 22 |
| J. Bourgain | 6919 | 44.9 | 42 | 6590 | 43.8 | 40 |
| H. Brezis | 11468 | 57.8 | 50 | 8386 | 49.5 | 48 |
| F. Browder | 2815 | 28.7 | 22 | 2807 | 28.6 | 22 |
| W. Browder | 646 | 13.7 | 13 | 547 | 12.6 | 12 |
| R. Bryant | 1489 | 20.8 | 21 | 1228 | 18.9 | 20 |
| L. Caffarelli | 6745 | 44.3 | 42 | 6280 | 42.8 | 41 |
| E. Calabi | 1224 | 18.9 | 18 | 1224 | 18.9 | 18 |
| L. Carleson | 1980 | 24 | 18 | 1484 | 20.8 | 17 |
| S-Y. Alice Chang | 1828 | 23.1 | 24 | 1806 | 22.9 | 24 |
| J. Cheeger | 3387 | 31.4 | 30 | 3348 | 31.2 | 30 |
| D. Christodoulou | 783 | 15.1 | 17 | 594 | 13.2 | 16 |
| A. Connes | 6475 | 43.5 | 43 | 5318 | 39.4 | 43 |
| I. Daubechies | 4674 | 36.9 | 28 | 3002 | 29.6 | 27 |
| P. Deift | 3004 | 29.6 | 26 | 2545 | 27.2 | 26 |
| P. Deligne | 6567 | 43.8 | 36 | 5592 | 40.4 | 33 |
| P. Diaconis | 3233 | 30.7 | 30 | 2970 | 29.4 | 30 |
| S. Donaldson | 2712 | 28.1 | 29 | 2277 | 25.8 | 29 |
| E. Dynkin | 1583 | 21.5 | 20 | 1090 | 17.8 | 16 |
| Y. Eliashberg | 1628 | 21.8 | 20 | 1460 | 20.6 | 18 |
| L. Faddeev | 1820 | 23 | 20 | 1285 | 19.4 | 18 |
| C. Fefferman | 3828 | 33.4 | 29 | 3815 | 33.4 | 29 |
| M. Freedman | 1207 | 18.8 | 16 | 990 | 17 | 16 |
| W. Fulton | 5890 | 41.4 | 27 | 1424 | 20.4 | 20 |
| H. Furstenberg | 2064 | 24.5 | 21 | 1650 | 21.9 | 21 |
| D. Gabai | 1314 | 19.6 | 17 | 1314 | 19.6 | 17 |
| J. Glimm | 1826 | 23.1 | 18 | 1419 | 20.3 | 18 |
| R. Graham | 3881 | 33.6 | 25 | 2280 | 25.8 | 24 |
| U. Grenander | 895 | 16.1 | 13 | 227 | 8.1 | 6 |
| P. Griffiths | 4581 | 36.5 | 26 | 1692 | 22.2 | 22 |
| M. Gromov | 7671 | 47.3 | 41 | 6200 | 42.5 | 38 |
| B. Gross | 1692 | 22.2 | 25 | 1635 | 21.8 | 24 |


| Member | $N_{\text {citations }}$ | Rule of thumb est. | $h$ | non-books only | revised est. | revised $h$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| V. Guillemin | 3710 | 32.9 | 27 | 2035 | 24.4 | 22 |
| R. Hamilton | 2490 | 26.9 | 20 | 2392 | 26.4 | 19 |
| M. Hochster | 1727 | 22.4 | 22 | 1657 | 22 | 21 |
| H. Hofer | 2140 | 25 | 25 | 1928 | 23.7 | 24 |
| MJ. Hopkins | 714 | 14.4 | 17 | 714 | 14.4 | 17 |
| R. Howe | 1680 | 22.1 | 22 | 1579 | 21.5 | 22 |
| H. Iwaniec | 2822 | 28.7 | 26 | 1825 | 23.1 | 24 |
| A. Jaffe | 794 | 15.2 | 9 | 277 | 9 | 8 |
| P. Jones | 1112 | 18 | 19 | 1112 | 18 | 19 |
| V. Jones | 2025 | 24.3 | 18 | 1669 | 22.1 | 17 |
| R. Kadison | 1922 | 23.7 | 20 | 1042 | 17.4 | 18 |
| R. Kalman | 558 | 12.8 | 10 | 428 | 11.2 | 10 |
| N. Katz | 2370 | 26.3 | 23 | 1582 | 21.5 | 20 |
| D. Kazhdan | 2332 | 26.1 | 27 | 2332 | 26.1 | 27 |
| R. Kirby | 963 | 16.8 | 15 | 678 | 14.1 | 14 |
| S. Klainerman | 2324 | 26 | 28 | 2144 | 25.0 | 27 |
| J. Kohn | 1231 | 18.9 | 19 | 1068 | 17.6 | 18 |
| J. Kollár | 3100 | 30.1 | 26 | 1947 | 23.8 | 22 |
| B. Kostant | 2509 | 27 | 25 | 2509 | 27 | 25 |
| R. Langlands | 1466 | 20.6 | 19 | 773 | 15.0 | 15 |
| H.B. Lawson | 2576 | 27.4 | 22 | 1846 | 23.2 | 21 |
| P. Lax | 4601 | 36.6 | 30 | 3560 | 32.2 | 27 |
| E. Lieb | 5147 | 38.7 | 33 | 4349 | 35.6 | 32 |
| T. Liggett | 1975 | 24 | 17 | 984 | 16.9 | 16 |
| L. Lovasz | 5638 | 40.5 | 34 | 4259 | 35.2 | 30 |
| G. Lusztig | 5786 | 41.1 | 40 | 4945 | 38.0 | 38 |
| R. MacPherson | 2031 | 24.3 | 22 | 1676 | 22.1 | 21 |
| G. Margulis | 2267 | 25.7 | 26 | 1788 | 22.8 | 25 |
| J. Mather | 1399 | 20.2 | 21 | 1399 | 20.2 | 21 |
| B. Mazur | 2842 | 28.8 | 26 | 2440 | 26.7 | 24 |
| D. McDuff | 2289 | 25.8 | 24 | 1417 | 20.3 | 23 |
| H. McKean | 2480 | 26.9 | 24 | 1701 | 22.3 | 21 |
| C. McMullen | 1738 | 22.5 | 25 | 1368 | 20.0 | 24 |
| J. Milnor | 7856 | 47.9 | 48 | 4559 | 36.5 | 40 |
| C. Morawetz | 789 | 15.1 | 16 | 764 | 14.9 | 15 |
| J. Morgan | 1985 | 24.1 | 25 | 1484 | 20.8 | 21 |
| G. Mostow | 1180 | 18.5 | 19 | 896 | 16.2 | 17 |
| J. Nash | 1337 | 19 | 10 | 1337 | 19 | 10 |
| E. Nelson | 1010 | 17.2 | 15 | 753 | 14.8 | 11 |
| L. Nirenberg | 9145 | 51.6 | 45 | 8781 | 50.6 | 43 |


| Member | $N_{\text {citations }}$ | Rule of thumb est. | $h$ | non-books only | revised est. | revised $h$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| S. Novikov | 2368 | 26.3 | 27 | 1677 | 22.1 | 21 |
| A. Okounkov | 1677 | 22.1 | 24 | 1677 | 22.1 | 24 |
| D. Ornstein | 1100 | 17.9 | 19 | 1022 | 17.3 | 18 |
| J. Palis | 1570 | 21.4 | 19 | 895 | 16.2 | 18 |
| P. Rabinowitz | 6633 | 44 | 29 | 5316 | 39.4 | 29 |
| M. Ratner | 506 | 12.1 | 13 | 506 | 12.1 | 13 |
| K. Ribet | 1022 | 17.3 | 18 | 1021 | 17.3 | 18 |
| P. Sarnak | 3114 | 30.1 | 32 | 2780 | 28.5 | 29 |
| M. Sato | 738 | 14.7 | 12 | 738 | 14.7 | 12 |
| R. Schoen | 3945 | 33.9 | 34 | 3493 | 31.9 | 34 |
| J. Serre | 10119 | 54.3 | 53 | 4481 | 36.1 | 36 |
| C. Seshadri | 984 | 16.9 | 15 | 831 | 15.6 | 14 |
| Y. Sinai | 3357 | 31.3 | 31 | 2547 | 27.3 | 28 |
| I. Singer | 2982 | 29.5 | 28 | 2951 | 29.3 | 28 |
| Y. Siu | 1494 | 20.9 | 22 | 1350 | 19.8 | 21 |
| S. Smale | 4581 | 36.5 | 39 | 3942 | 33.9 | 36 |
| R. Solovay | 781 | 15.1 | 14 | 781 | 15.1 | 14 |
| J. Spencer | 758 | 14.9 | 15 | 1334 | 19.7 | 18 |
| R. Stanley | 6510 | 43.6 | 35 | 3148 | 30.3 | 32 |
| H. Stark | 678 | 14.1 | 13 | 653 | 13.8 | 13 |
| C. Stein | 763 | 14.9 | 12 | 658 | 13.9 | 11 |
| E. Stein | 14049 | 64 | 49 | 5788 | 41.1 | 37 |
| R. Steinberg | 1850 | 23.2 | 19 | 1068 | 17.6 | 17 |
| S. Sternberg | 2438 | 26.7 | 25 | 1476 | 20.8 | 18 |
| D. Stroock | 3299 | 31.0 | 27 | 2028 | 24.3 | 24 |
| D. Sullivan | 3307 | 31.1 | 32 | 3248 | 30.8 | 31 |
| R. Swan | 1109 | 18 | 20 | 998 | 17.1 | 19 |
| E. Szemerédi | 2536 | 27.2 | 26 | 2536 | 27.2 | 26 |
| T. Tao | 6730 | 44.3 | 40 | 6214 | 42.3 | 39 |
| J. Tate | 2979 | 29.5 | 30 | 2612 | 27.6 | 28 |
| C. Taubes | 1866 | 23.2 | 24 | 1626 | 21.8 | 23 |
| J. Thompson | 789 | 15.2 | 14 | 789 | 15.2 | 14 |
| J. Tits | 3463 | 31.8 | 28 | 2958 | 29.4 | 26 |
| K. Uhlenbeck | 1852 | 23.2 | 21 | 1756 | 22.6 | 20 |
| S. Varadhan | 2894 | 29 | 28 | 2153 | 25.1 | 26 |
| D. Voiculescu | 2952 | 29.3 | 29 | 2387 | 26.4 | 26 |
| A. Wiles | 1387 | 20 | 14 | 1387 | 20 | 14 |
| S-T. Yau | 7536 | 46.9 | 44 | 7066 | 45.4 | 43 |
| E. Zelmanov | 1055 | 17.5 | 16 | 1020 | 17.2 | 16 |


[^0]:    ${ }^{1}$ Citations pre-2000 in Mathscinet are not complete. Google scholar and Thompson Reuters' Web of Science also have sources of error. We decided that Mathscinet was our most complete option for analyzing mathematicians. For relatively recent Fields medalists, the effect of lost citations is reduced.

[^1]:    ${ }^{2}$ Actually, our computation of $\operatorname{Par}\left(N_{\text {citations }}\right)$ using generating series became not so easy on a desktop machine when $N_{\text {citations }}$ is a few thousand. Instead, one could use the Hardy-Ramanujun approximation $\operatorname{Par}\left(N_{\text {citations }}\right) \sim \frac{1}{4 N_{\text {citations }}^{3}} e^{\pi \sqrt{\frac{2 N_{\text {citations }}}{3}}}$. Even more precisely, one can use Wolfram Alpha, which gives the partition numbers for up to a million, which is well beyond our needs.

